

# Russia-Ukraine war and G7 debt markets: Evidence from public sentiment towards economic sanctions during the conflict

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## Abstract

War-related expectations cause changes to investors' risks and returns preferences. In this study, we examine the implications of war and sanctions sentiment for the G7 countries' debt markets during the Russia-Ukraine war. We use behavioural indicators across social media, news media, and internet attention to reflect the public sentiment from 1st January 2022 to 20th April 2023. We apply the quantile-on-quantile regression (QQR) and rolling window wavelet correlation (RWWC) methods. The quantile-on-quantile regression results show heterogeneous impact on fixed income securities. Specifically, extreme public sentiment has a negative impact on G7 fixed income securities return. The wavelets correlation result shows dynamic correlation pattern among public sentiment and fixed income securities. There is a negative relationship between public sentiment and G7 fixed income securities. The correlation is time-varying and highly event dependent. Our additional analysis using corporate bond data indicates the robustness of our findings. Furthermore, the contagion analysis shows public sentiment significantly influence G7 fixed income securities spillover. Our findings can be of great significance while framing strategies for asset allocation, portfolio performance and risk hedging.

## KEYWORDS

economic sanctions, fixed income securities, G7 debt, quantile approaches, Russia-Ukraine war

## 1 | INTRODUCTION

War events have long been recognised to have implications for financial markets (Frey & Kucher, 2001; Hudson & Urquhart, 2015; Schneider & Troeger, 2006). In essence, war-related expectations cause changes to investors' risks and returns preferences. This suggests that war related

events are the most important specific event that move the financial market in a significant way (e.g., Boubaker et al., 2022; Bounou & Yatié, 2022; Zaremba et al., 2022). More specifically, Frey and Kucher (2001) find that war events are priced in government bonds. Hudson and Urquhart (2015), on the other hand, evidenced limited linkages between war events and stock market returns in

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the United Kingdom. From a global financial market perspective while accounting for multiple war events, Schneider and Troeger (2006) evidenced significant negative stock market reactions. Typically, earlier studies on the impact of war events on financial markets are limited to what transpired between markets during the war event (see e.g., Izzeldin et al., 2023; Karkowska & Urjasz, 2023; Kumari et al., 2023; Lo et al., 2022). A recent strand in the literature uses behavioural indicators (i.e., media coverage, fake news, panic, sentiment, and media hype) to contextualise investor sentiments (Boungou & Yatié, 2022; Huynh et al., 2021; Khalfaoui et al., 2023) and thus, provides a direct measure of investor sentiments towards extreme events; however, in the framework of investor sentiments to war and sanctions, these behavioural indicators have been used in isolation to gauge investor sentiment.<sup>1</sup> Again, prior studies provide no understanding of how the financial market in general reacts to positive and negative investor sentiment towards economic sanction during the war in Ukraine.

This paper explores the implication of war and economic sanctions sentiments from a multidimensional perspective using a recently developed public sentiment index (RUWESent) constructed by combining Twitter Sentiments, Google Trend, Wikipedia Trend, and News Sentiments (Abakah et al., 2022). By this, we can account for public sentiment as an additional variable in the network analysis to provide a direct role of public sentiment in extreme events. Specifically, we contribute to this literature by examining the impact of war- and sanctions-induced uncertainty on the fixed-income markets in the G7 countries. We focus on debt markets because globally, fixed-income markets are the largest subgroup of financial markets by issuances and market capitalization. Fixed-income markets are three times larger than global equities markets making it a systematically important market with implications for the global economy.<sup>2</sup> Nonetheless, the growing body of literature on the impact of the Russia-Ukraine war presents sparse evidence on the possibility of a relationship between war- and sanctions-induced public sentiment and fixed income securities under varying market conditions (see e.g., Boubaker et al., 2022; Boungou & Yatié, 2022; Zaremba et al., 2022). Additionally, the time-frequency aspect of such a relationship is also missing. In this regard, two complimentary theoretical positions are apparent on the importance of investor sentiments on financial market dynamics. The first theoretical position is broadly consistent with the view that investors make financial decisions based on sentiments, which suggest that investor sentiments affect financial market dynamics (Huynh et al., 2021; Liang et al., 2020) in line with the theoretical model of noise

traders (Long et al., 1990). The second theoretical position is consistent with the heterogeneous market hypothesis (Müller et al., 1993), which suggests that war and sanction related sentiments may alter investors' risk and return preferences across bullish, bearish, and normal market conditions.

A key question that needs to be addressed is, why the G7 fixed income market is particularly sensitive to the ongoing war and the possible channels through which it could be influenced by the war in Ukraine. First, we note that the G7 countries are at the forefront of the sanctions regime that has been imposed on Russia as well as the financial and military support Ukraine is receiving in its fight with Russia. Consequently, these governments have increased spending on defence and other war-related expenses.<sup>3</sup> This can result in higher levels of debt and inflation, which can contribute to an increase in interest rates.<sup>4</sup> In turn, this can influence the pricing of interest-rate-sensitive fixed-income securities such as bonds. Second, war-related events can also make investors more risk-averse (Verdickt, 2020; Wang & Young, 2020). In uncertain times, investors may pursue safer investments (Costantini & Sousa, 2022; Mohamad, 2022), such as government bonds, resulting in a rise in demand for fixed income securities. This can result in higher bond prices and lower yields (Leippold & Matthys, 2022; Zaremba et al., 2022). On the possible channels, the Ukraine crisis may raise geopolitical risk and uncertainty, causing a fixed income market flight to safety (see e.g., Feng et al., 2023). This may stimulate demand for safe-haven assets like US Treasuries, lowering yields. Moreover, the global supply networks have been disrupted causing increased inflation, supply chain costs, and central bank interest rates. Consequently, investors in fixed-income securities may sell off.

Our paper contributes to literature in four main ways. First, we contribute to the literature that examines the impact of war on the debt markets (Frey & Kucher, 2001). We focus on war- and sanctions-induced sentiments and show a negative effect on fixed income securities return in both bullish and bearish markets; however, when the range of sentiment is considered normal, the effect is positive, suggesting an asymmetric static impact of war and sanctions sentiment on fixed income securities. This finding supports the heterogeneous market hypothesis (Müller et al., 1993), and suggests that war- and sanction-related sentiments may alter investors' risk and return preferences across bullish, bearish, and normal market conditions. Second, we extend the literature that examines the impact of war- and sanctions-induced uncertainty on financial market (Boubaker et al., 2022; Boungou & Yatié, 2022; Zaremba et al., 2022). The rolling window wavelet correlation is used to provide evidence

on the dynamics of the relationship between war- and sanctions-related sentiments and debt markets. We find a strong negative correlation between investor sentiments and fixed income securities returns at the start of the invasion and during major invasion events, implying that war- and sanctions-induced investor sentiments had a time-varying influence. As a result, negative sentiment has a negative impact on fixed income securities, indicating that it causes irrational investor behaviour as well as increase noise trader loss aversion and herding behaviour (Long et al., 1990). Furthermore, the negative correlation in the higher decomposition level during the financial turbulence period suggests that returns on fixed income securities assets would offer hedging opportunities with other safe-haven assets and diversification opportunities for short- and long-term investors (Baur & Lucey, 2010). Third, we contribute to the literature studying the contagion effect of the Russia-Ukraine war on the global financial markets (see e.g., Bounou & Yatié, 2022; Khalfaoui et al., 2023). We demonstrate that the fixed income markets of Japan, United Kingdom, United States, and UK are net risk receivers, whereas Italy, Germany France are the major transmitter of risk suggesting that there is contagion effect from the Eurozone's to other G7 countries during Russia Ukraine war. Lastly, we provide novel evidence on the subset of the public sentiment towards war and sanction. This includes the Russia-Ukraine war sentiment index, Russia-Ukraine War Sanction sentiment index, and Russia-Ukraine War anxiety index. The results indicate the effect is stronger for Russia-Ukraine War sentiment index and the highest negative impact is witnessed on the fixed income market in Germany.

The remainder of the study is structured as follows. Section 2 provides a brief overview of the literature. Section 3 reports the data and methodology, and section 4 discusses the results. Finally, section 5 concludes the paper.

## 2 | BRIEF OVERVIEW OF THE LITERATURE

Our paper is related to recent strand of the literature on the impact of the Russia-Ukraine war sentiments on financial markets. These studies reveal the potential impact of war events on the financial markets. Earlier studies in this literature widely acknowledge that war events have a significant effect on the interactions at the core of financial markets around the world (Ahmed et al., 2023; Bounou & Yatié, 2022; Choudhry, 2010; Frey & Kucher, 2000, 2001; Hudson & Urquhart, 2015; Schneider & Troeger, 2006). Since the outbreak of the Russia-Ukraine war, a few studies have examined

the impact of the conflict on markets (Ahmed et al., 2023; Boubaker et al., 2022; Bounou & Yatié, 2022; Chortane & Pandey, 2022; Costola & Lorusso, 2022; Halousková et al., 2022; Kumari et al., 2022; Tosun & Eshraghi, 2022; Umar, Riaz, & Yousaf, 2022; Yousaf et al., 2022), but none of them examined the impact of the conflict on the debt market. This strand of the literature examines cryptocurrency assets in the context of a network of traditional financial assets (Adekoya et al., 2022; Diaconasu et al., 2022; Mohamad, 2022; Umar, Polat, et al., 2022); and focus on what transpired between markets during the period. It should also be noted that the Caldara and Iacoviello (2022) geopolitical risk (GPR) index, Google Trends, and news events have been used as the principal proxies for sentiments during the Russia-Ukraine war (see Aslanidis et al., 2022; Będowska-Sójka et al., 2022; Halousková et al., 2022; Long et al., 2022; Qureshi et al., 2022).

Regarding the implications of the war on stock markets, Kumari et al. (2022) demonstrated that in the case of the Russia-Ukraine war (the invasion), the stock markets of the closest European countries (Poland, Denmark, and Portugal) to the war region experienced event day impact, but recovered in the post-event period, showing positive cumulative abnormal returns (CARs). Furthermore, Boubaker et al. (2022) demonstrated that the invasion produced CARs for global stock market indices, albeit with varying effects. They specifically evidenced that markets in NATO countries had higher returns. The positive outcomes observed after these countries' events indicate an improvement in market sentiment as investors regained confidence that NATO would not pursue armed conflict. Nevertheless, the sanctions had adverse effects on stock returns, signalling that investors anticipated trade disruptions and a decline in the value of firms. Additionally, attention biases caused by salience resulted in investors favouring stocks with considerable growth potential while undervaluing stocks with significant drawbacks. As a result, higher returns were observed in the subsequent period (Umar, Riaz, & Yousaf, 2022).

Additional research has explored the impact of the invasion on various markets, including clean energy, conventional energy, and metals. These studies have shown that investors have given greater importance to the likelihood of outcomes associated with the clean energy sector due to the war's developments. As a result, the objective probability of outcomes has been distorted (Umar, Riaz, & Yousaf, 2022). Adekoya et al. (2022) examined how oil markets related to bonds, Bitcoin, the US dollar, gold, and stocks before and during the Russia-Ukraine war using intraday data. They reported that the link between oil and other financial assets is stronger than before the war. Oil, rather than being a net receiver of spillovers, becomes a net transmitter of spillovers during the war.

The net directional pairwise results show that oil connects to the remaining assets in a variety of ways before the war, but it has a large spillover effect on all of them during the war. They also demonstrate that the spillover effect diminishes over time.

Another body of literature has focused on examining the hedging characteristics of various asset classes within the context of the Russian-Ukraine conflict. These studies have provided insights into effective strategies for mitigating geopolitical risk. Notably, different asset classes have demonstrated varying degrees of sensitivity to risk, both in terms of magnitude and timescale. Bonds and stocks, for instance, have exhibited strong correlation over multi-week periods, while currencies have been affected over shorter timescales (Będowska-Sójka et al., 2022). Finally, Qureshi et al. (2022) compiled a database of news events related to the conflict and explored the systemic risk implications on financial sector indices. The events were categorised as political actions, military actions, sanctions, military support for Ukraine, financial support for Ukraine, and Russian invasion. They demonstrated that the conflict's systemic instability costs extend beyond Russia and Ukraine and that, sanctions have a systemic risk spillover effect on European countries and the United States.

In this study, we investigate the implications of war and economic sanctions sentiment for the G7 debt market using a sentiment index developed to reflect the public sentiment across social media, news media, and internet attention.

### 3 | DATA AND METHODOLOGY

#### 3.1 | Data

The war between Russia and Ukraine is the result of long-standing tensions between the two countries, which began in late December 2021 and culminated in Russia's invasion of Ukraine on 24 February 2022. To represent the proxy of Russia-Ukraine war induced economic sanctions sentiments we use novel measure of Russia-Ukraine War Economic Sanctions News Sentiment Index (RUWESsent) which is developed based on Twitter Sentiments (TS), Google Trend (GT), Wikipedia Trend (WT), and News Sentiments (NS). In RUWESsent, TS represents public sentiment towards war and sanctions, GT and WT indicate anxiety (stress) or intensity due to intense interest in war and sanctions, and NS represents media sentiment towards war and sanctions (see Abakah et al., 2022). To avoid biases, we chose a study sample period from 1 January 2022

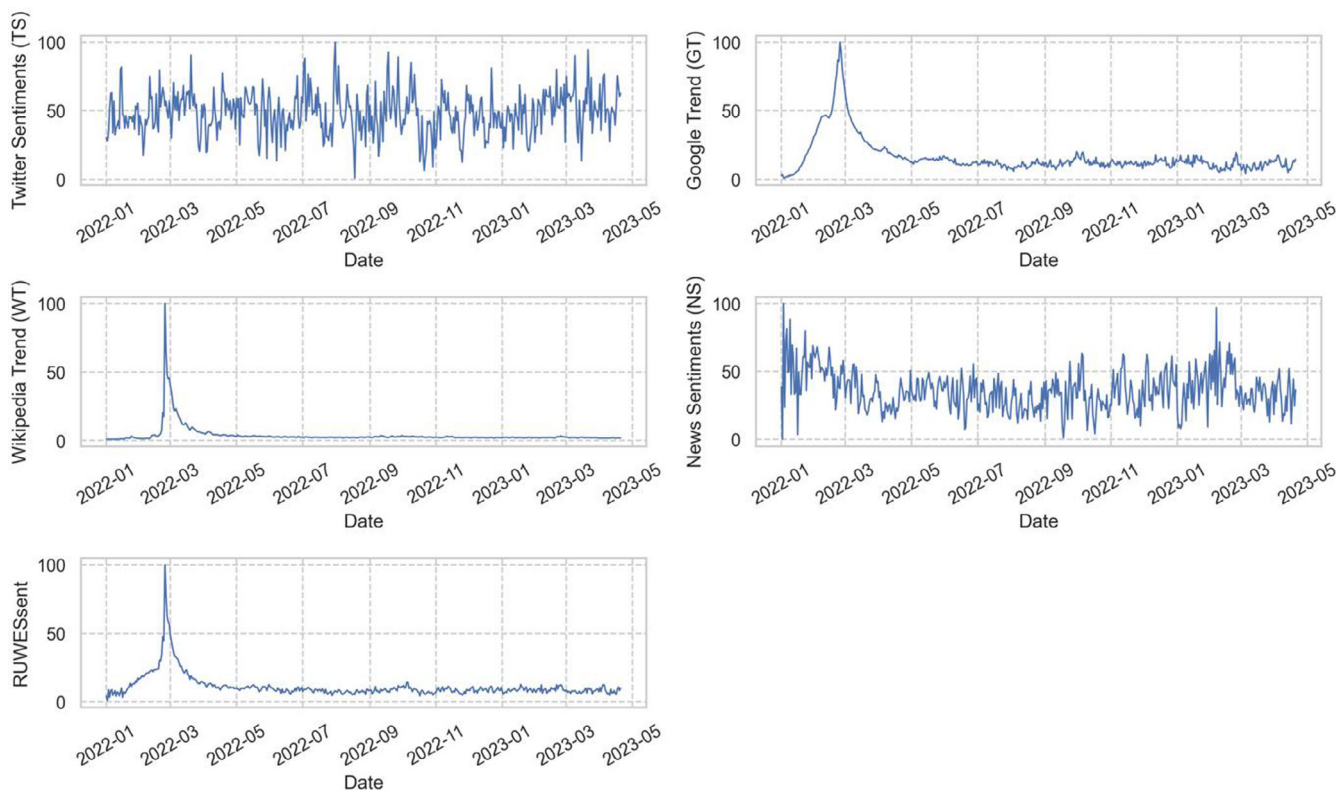


FIGURE 1 RUWESsent and its indicators. Source authors illustration from the data of RUWESsent website <https://ruwessent.wordpress.com/> [Colour figure can be viewed at [wileyonlinelibrary.com](https://wileyonlinelibrary.com)]

to 20 April 2023, where the ending period is based on the availability of RUWESsent. The value of RUWESsent is ranging from 1 to 100, with a low value (1–49) representing negative sentiment and a high value (50–100) representing positive sentiment. The indicators of RUWESsent are illustrated in Figure 1, which shows RUWESsent was at its peak

TABLE 1 List of variables.

Variables	Description	Data source
Canada	S&P Canada Sovereign Bond Index	Datastream
France	S&P France Sovereign Bond Index	Datastream
Germany	S&P Germany Sovereign Bond Index	Datastream
Italy	S&P Italy Sovereign Bond Index	Datastream
Japan	S&P Japan Sovereign Bond Index	Datastream
UK	S&P U.K. Gilt Index	Datastream
US	S&P U.S. Treasury Bond Index	Datastream
RUWESsent	Russia-Ukraine War induced economic sanctions sentiments index	<a href="https://ruwessent.wordpress.com/">https://ruwessent.wordpress.com/</a>

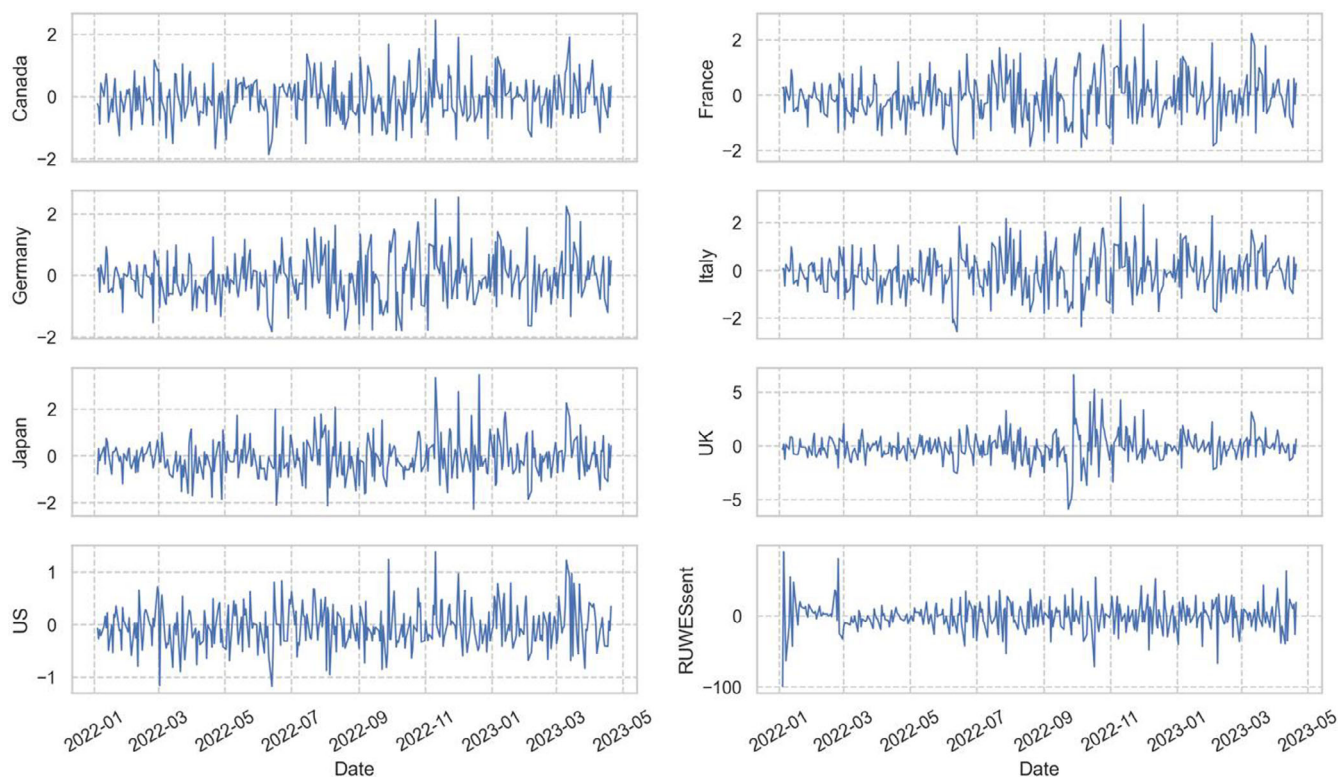


FIGURE 2 Historical series of daily percentage changes in variables. [Colour figure can be viewed at [wileyonlinelibrary.com](http://wileyonlinelibrary.com)]

on 24 February 2022. Finally, for debt market, we use daily data of treasury/sovereign/quasi-government bond indices of G7 countries.<sup>5</sup> We chose G7 countries because they are at the forefront of the sanctions regime that has been imposed on Russia as well as the financial and military support Ukraine is receiving in its fight with Russia. Table 1 contains a summary of the selected securities and RUWESsent with data sources.

As the level series of variables are non-stationary according to Augmented Dickey–Fuller (ADF) and Phillips–Perron (PP) unit-root test, we transformed the series using percentage change as  $r_{i,t} = 100 \times (P_{i,t} - P_{i,t-1}) / (P_{i,t-1})$ , which are presented in Figure 2. We can see similar return movement of selected assets in some specific periods. Table 2 presents the summary statistics for all the series. In Panel A, the fixed income securities exhibit a negative mean return, indicating a negative return throughout the sample period. On the other hand, the mean return of RUWESsent is positive. Additionally, all the series demonstrate positive skewness, indicating a distribution with a longer tail on the right side. The ADF and PP unit root tests provide strong evidence supporting the stationarity process for all the series. Finally, we examine the Pearson correlation and results are presented in Panel B, which shows a positive correlation with all fixed income security indices and RUWESsent.

TABLE 2 Summary statistics and correlation analysis results.

	Canada	France	Germany	Italy	Japan	UK	US	RUWESsent
<b>Panel A: Summary statistics</b>								
<i>N</i>	338	338	338	338	338	338	338	338
Mean	−0.04	−0.07	−0.06	−0.06	−0.06	−0.11	−0.03	2.27
Std.	0.66	0.81	0.77	0.87	0.83	1.33	0.4	22.43
Minimum	−1.84	−2.12	−1.81	−2.53	−2.25	−5.73	−1.17	−62.95
Maximum	2.49	2.76	2.57	3.12	3.54	6.86	1.4	146.32
Skewness	0.24	0.31	0.32	0.27	0.69	0.6	0.31	1.64
Kurtosis	0.53	0.41	0.43	0.57	2.11	4.62	0.48	7.87
ADF	−6.9***	−6.927***	−7.171***	−7.218***	−6.226***	−6.936***	−6.783***	−7.059***
PP	−316.698***	−320.707***	−325.782***	−307.263***	−340.791***	−318.816***	−324.643***	−404.497***
<b>Panel B: Correlation analysis</b>								
Canada	1							
France	0.97***	1						
Germany	0.97***	1.00***	1					
Italy	0.95***	0.99***	0.99***	1				
Japan	0.95***	0.99***	0.99***	0.98***	1			
UK	0.98***	0.98***	0.98***	0.96***	0.96***	1		
US	0.97***	0.97***	0.97***	0.95***	0.97***	0.96***	1	
RUWESsent	0.44***	0.48***	0.48***	0.47***	0.51***	0.44***	0.47***	1

Note: *N*, number of observations; Std., standard deviation; ADF, augmented Dickey–Fuller unit-root test; PP, Phillips–Perron unit-root test; \*\*\*, \*\*, \* denote significance at 1%, 5%, and 10% significance level.

### 3.2 | Methods

This paper uses two robust estimation techniques to examine the effect, and dynamic correlation between the RUWESsent on fixed income securities market: (1) Quantile on Quantile Regression (Sim & Zhou, 2015) and (2) Rolling window wavelet correlation approach (Polanco-Martinez et al., 2018). Moreover, we also examine the robustness of our results using quantile granger causality test based on Troster (2018) and examined contagion across G7 debt market using time-varying parameter vector autoregressions (TVP-VAR) based connectedness approach based on Antonakakis et al. (2020).<sup>6</sup>

#### 3.2.1 | Quantile on quantile regression

First, we use the quantile on quantile regression (QQR) method proposed by Sim and Zhou (2015) to examine the overall impact of the RUWESsent on fixed income securities market. The QQR is a more advanced version of simple quantile regression (QR) that is created by combining non-parametric estimations with basic QR. The traditional QR model can investigate the effect of the independent

variable on different quantiles and the dependent variable's conditional mean (Sim & Zhou, 2015). The QQR model, on the other hand, combines traditional linear regression with the QR model, which can analyse the impact of both dependent and independent variables over conditional distributions and provides a better understanding of the interactions between the dependent and independent variables. While QR can show how the RUWESsent have affected different quantiles of fixed income securities returns, we cannot see how the numerous different quantiles of RUWESsent have affected fixed income securities returns. The QQR contributes to this by regressing different quantiles of RUWESsent on different quantiles of fixed income securities returns. This is necessary because, depending on the market condition (bear, normal, or bull), higher and lower sentiments have different magnitude effects on assets return (Bossman et al., 2023). Sim and Zhou (2015) proposed the non-parametric function of the QQR model as:

$$r_t = \beta^{\theta} RUWESsent_{t-1} + \alpha^{\theta} r_{t-1} + \varepsilon_t^{\theta} \quad (1)$$

where  $r_t$  denotes daily returns of fixed income securities and  $RUWESsent_t$  denotes changes in the sentiments at

the time  $t$ . Moreover,  $\theta$  denotes fixed income securities returns at  $\theta_{th}$  quantile and  $\beta^\theta(\cdot)$  is unknown.  $RUWESsent^\tau$  is calculated first-order Taylor approximation using the following equation:

$$\begin{aligned} \beta^\theta RUWESsent_{t-1} &\approx \beta^\theta RUWESsent^\tau \\ &+ \beta^\theta (RUWESsent^\tau) (RUWESsent_{t-1} - RUWESsent^\tau) \\ &\equiv b_0(\theta, \tau) + b_1'(\theta, \tau) \times (RUWESsent_{t-1} - RUWESsent^\tau) \end{aligned} \quad (2)$$

where  $\tau$  represents the  $\tau_{th}$  quantile of Russia-Ukraine war and sanctions sentiments. Afterward, the substitution of first-order Taylor approximation in Equation (1) with Equation (2) outputs Equation (3).

$$\begin{aligned} r_t &= \beta^\theta (RUWESsent^\tau) \\ &+ \beta^\theta (RUWESsent^\tau) (RUWESsent_{t-1} - RUWESsent^\tau) \\ &+ \alpha^\theta r_{t-1} + \varepsilon_t^\theta \end{aligned} \quad (3)$$

where Equation (3) is explained by following equation:

$$\begin{aligned} \min_{b_0(\theta, \tau), b_1(\theta, \tau), \alpha^\theta(\tau)} &\sum_{t=1}^T \rho_\theta \left[ r_t - b_0 \right. \\ &- b_1 (RUWESsent_{t-1} - RUWESsent^\tau) \\ &\left. - \alpha^\theta r_{t-1} \right] K \left( \frac{RUWESsent(RUWESsent_{t-1}) - \tau}{h} \right) \end{aligned} \quad (4)$$

where  $\rho_\theta(\cdot)$  solved  $\theta$ th quantile of  $r_t$  represents the absolute value of the slope function and the Gaussian kernel function  $K(\cdot)$  with  $h$  is the bandwidth. Here,  $h$  is selected using the cross-validation method.

### 3.2.2 | Rolling window wavelet correlation

The application of the QQR can reveal salient information about how the  $RUWESsent$  impacts different quantiles of fixed income securities returns. However, the effect of the  $RUWESsent$  might be time-varying. Therefore, we apply rolling window wavelet correlation (RWWC) proposed by Polanco-Martínez et al. (2018), which is elaborated as follows.

First, based on the work of Polanco-Martínez et al. (2018), we compute maximal overlap discrete wavelet transform (MODWT) wavelet correlation, to examine the impact Russia-Ukraine war and sanctions induced sentiments on fixed income securities returns at different time periods. The equitable wavelet correlation for scale  $\lambda_j$  between two time series  $X$  and  $Y$  is presented as follows:

$$\tilde{\rho}_{XY} = \frac{\text{cov}(\tilde{W}_{Yjt}, \tilde{W}_{Xjt})}{\sqrt{\text{var}\{\tilde{W}_{Xjt}\}\text{var}\{\tilde{W}_{Yjt}\}}} = \frac{\tilde{\gamma}_{XY}(\lambda_j)}{\tilde{\sigma}_X^2(\lambda_j)\tilde{\sigma}_Y^2(\lambda_j)} \quad (5)$$

In the preceding expression,  $\tilde{\gamma}_{XY}(\lambda_j)$  signifies the equitable approximator of wavelet covariance between market coefficient  $\tilde{W}_{Yjt}$  and  $\tilde{W}_{Xjt}$ , whereas  $\tilde{\sigma}_X^2(\lambda_j)$  and  $\tilde{\sigma}_Y^2(\lambda_j)$  are the unbiased estimators of wavelet variances  $X$  and  $Y$ , respectively, associated with the scale  $\lambda_j$ . MODWT defines as follows:

$$\tilde{\sigma}_X^2(\lambda_j) = \frac{1}{N_j} \sum_{t=L_{j-1}}^{N-1} \tilde{W}_{j,t}^2 \quad (6)$$

where  $\tilde{W}_{j,t}^2$  signifies  $j$ th scale of MODWT coefficient for  $X$ ,  $L_j = (2j-1)(L-1)+1$  signifies length of scale  $\lambda_j$ . Following Daubechies (1992), we construct confidence interval  $100(1-2p)\%$  for wavelet coherence. An expression for  $100(1-2p)\%$  confidence interval for wavelet coherence is given by  $\tan h\{h[\tilde{\rho}_{XY}(\lambda_j)]\phi^{-1}(i-p)/\sqrt{N_j-3}\}$ , where  $\phi^{-1}(p)$  represents  $100p\%$  points for standard normal distribution and  $h(\tilde{\rho}_{XY}) = \tan h^{-1}(\tilde{\rho}_{XY})$  denotes the Fisher Z-transformation (Daubechies, 1992; Gençay et al., 2001).

We employ a dynamic rolling window wavelet correlation method to assess correlation across multiple dimensions in both time and frequency domains, allowing us to analyse how the correlation changes over time. This measure has proven to be highly advantageous in various finance studies due to its ability to analyse different time intervals effectively (Polanco-Martínez et al., 2018; Rehman, 2020). In our analysis, we set two decomposition levels ( $j=2$ ) and a 22-day rolling window per month, moving one data point at a time and centering it around a specific time.

## 4 | RESULTS

### 4.1 | QQR results

The results of the QQR method are presented in Figure 3. The graphs show the slope coefficients  $\beta_1(\theta, \tau)$ , which represent the impact of the  $\tau_{th}$  quantile of  $RUWESsent$  on fixed income securities returns, where the high and low quantiles of  $\theta_{th}$  represent the bullish and bearish fixed income securities market, respectively. Furthermore, because of the non-parametric process used in QQR estimations, it is not practical to examine the

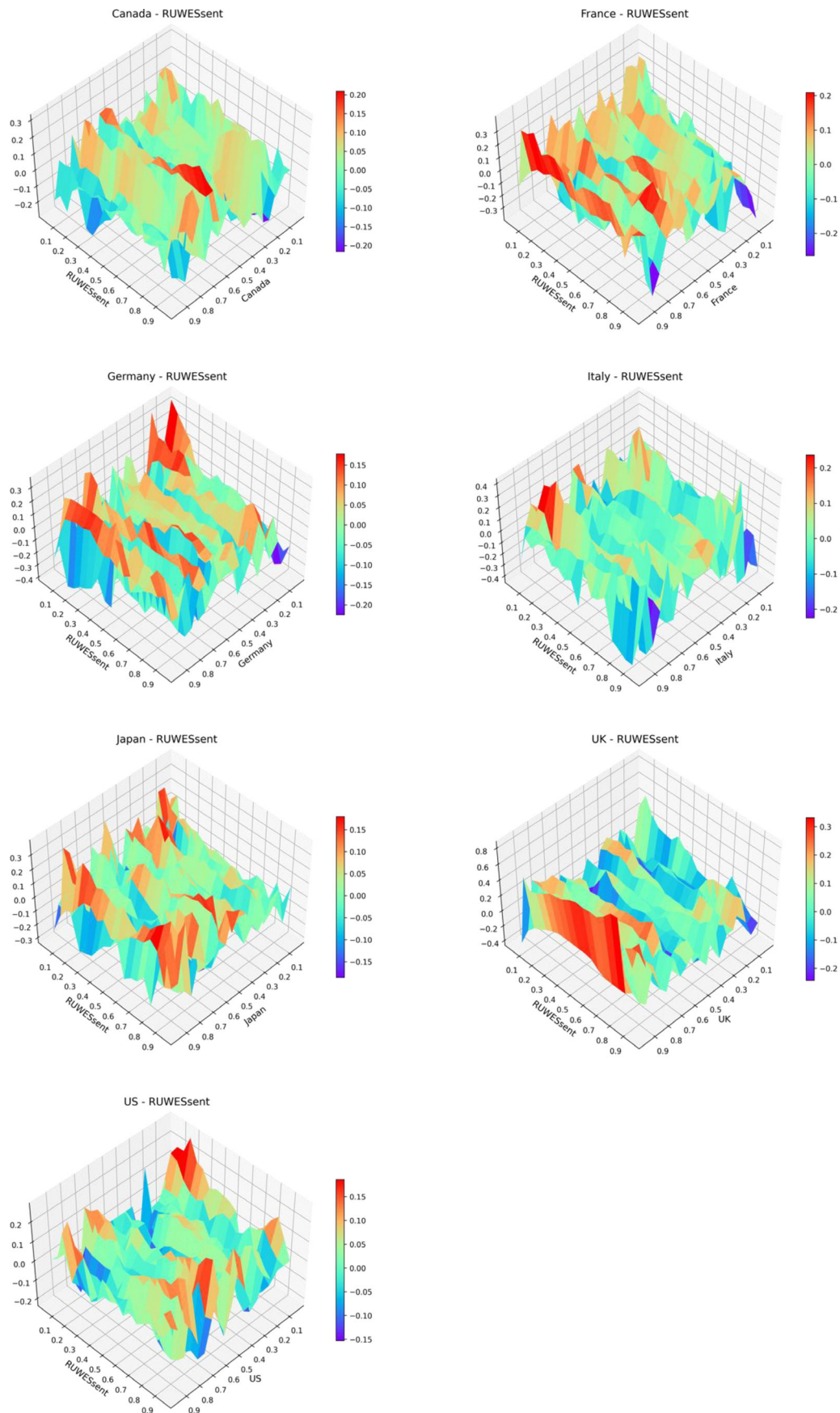


FIGURE 3 Quantile-on-quantile regression results between RUWESsent and US debt market. [Colour figure can be viewed at [wileyonlinelibrary.com](http://wileyonlinelibrary.com)]

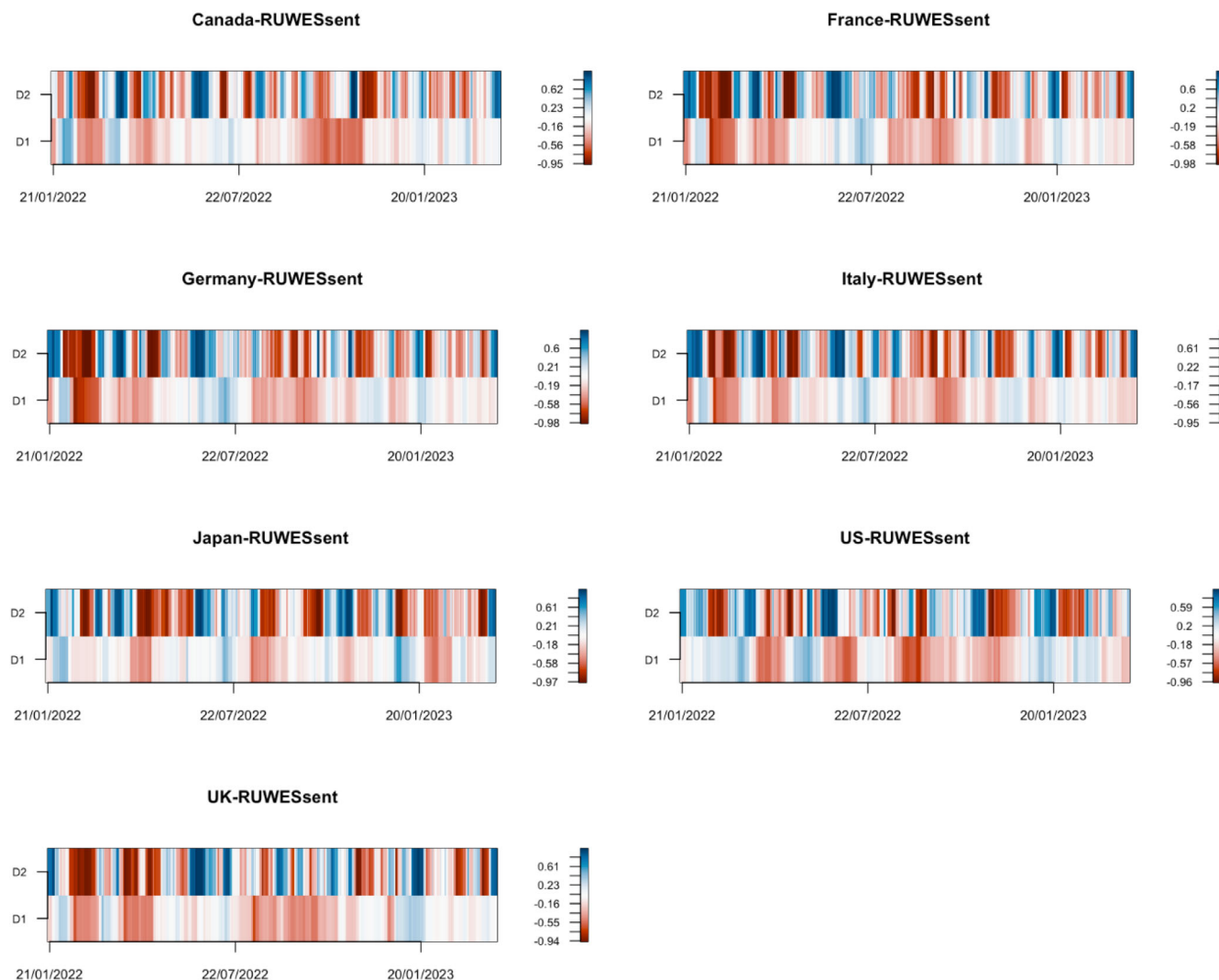


FIGURE 4 Rolling wavelet window correlation result results between RUWESsent and US debt market. [Colour figure can be viewed at [wileyonlinelibrary.com](http://wileyonlinelibrary.com)]

significance levels of the coefficients (Bossman et al., 2023; Sim & Zhou, 2015). According to Figure 4, each quantile of the RUWESsent has a different effect on fixed income securities returns across quantiles, which suggests a heterogeneous effect across quantiles. For example, in the case of United States, there is a negative relationship when both RUWESsent and United States are in the upper quantile (0.85–0.95), whereas there is a positive effect when both are in lower quantile (0.05–0.35). The results for other G7 countries fixed income securities show a similar heterogeneous effect across quantiles, for example, Canada, France, Germany, Italy, and United Kingdom. In contrast, the results for Japan show a positive impact in both extreme lower and higher quantiles and heterogeneous impact across rest of the quantiles.

The robustness of our QQR results is examined by comparing with the coefficient of the quantile regression

(QR) following Bossman et al. (2023), and the results are presented in Figure A1. The graph shows the mean coefficient of each quantile of QQR and coefficient of QR. As the QQR results do not deviate significantly from the QR, this indicates the robustness of QQR results. Additionally, supplementary to the QRR estimates, we adopted the quantile Granger causality test for further robustness and present the results in Tables A1 and A2. The finding shows there is unidirectional causal relationship with RUWESsent and G7 fixed income securities. This result indicates war and sanction related sentiments granger cause fixed income securities, but fixed income securities does not granger cause sentiments. This result is consistent with earlier studies by Zhang et al. (2018) who documented sentiments can predict asset price movement.

Overall, the QQR result shows that the upper extreme quantiles (0.05 and 0.95) of RUWESsent have a negative

effect on G7 fixed income securities return in a bullish market and bearish market, except Japan. Our findings indicate European, United Kingdom, and United States has considerable negative shock from Russia-Ukraine war. However, there is a positive impact on fixed income securities when sentiment is considered normal (0.35–0.65). As a result, our findings support the heterogeneous market hypothesis (Müller et al., 1993), and suggests that investors consider all available information, including the RUWESsent, when analysing their risk and return preferences across bullish, bearish, and normal market conditions (Bossman et al., 2023). However, the above QQR method results show the asymmetric static impact of the RUWESsent on fixed income securities, which fails to capture the time-varying relationship. As a result, in the following section, we investigate the time-frequency aspect of the relationship between war and sanctions-induced investor sentiment and the fixed-income market using the dynamic rolling window wavelet correlation.

## 4.2 | The RWWC results

The results of RWWC are illustrated in Figure 4, where values are usually between negative and positive, indicating a negative and positive correlations. The RWWC shows some intriguing results that cannot be obtained with the QQR method. For example, the relationship's strength varies over time and is determined by how frequently changes in the RUWESsent coincide with fixed income securities returns. Correlation coefficients or time horizons are implied with changes of 1–5 days and intra-week to monthly periods for the two wavelet scales, or from D1 to D2. Higher scales, such as D2, describe processes that occur at a lower frequency, such as economic uncertainty, monetary policy, trade, and common shocks (Polanco-Martínez et al., 2018). This implies that volatility events are more closely associated with the first wavelet scales. The result shows France, Germany, and United Kingdom have high negative correlations, followed by Canada, and Italy to a lesser extent, and United States, and Japan have low negative correlations with the RUWESsent. The findings of QQR also shows similar pattern, where Japan is less affected by war and sanction sentiment. The dynamic result shows high negative correlation between RUWESsent and fixed income securities returns during the commence of the invasion and major invasion events. Furthermore, the results show a higher negative correlation in September 2022.

Therefore, the RWWC findings show a stronger negative link immediately following the invasion and

during September 2022, demonstrating that the RUWESsent index had a time-varying influence. As a result, negative sentiment has a negative impact on fixed income securities, signalling that negative sentiment causes irrational investor behaviour as well as increasing noise trader loss aversion and herding behaviour (Long et al., 1990). Furthermore, the negative correlation in the higher decomposition level during the financial turbulence period suggests that returns on fixed income securities assets would offer hedging opportunities with other safe-haven assets and diversification opportunities for short- and long-term investors (Baur & Lucey, 2010).

Overall, our results of QQR and RWWC could be explained by the ongoing Russia-Ukraine war and economic sanctions, which have increased investors' loss aversion and herding behaviour. Earlier studies on other assets also uncover the negative impact of Russia Ukraine war (e.g., Bounou & Yatié, 2022; Yousaf et al., 2022). Moreover, previous research also found that negative sentiments increase loss aversion and herding behaviour and have a negative impact on the other assets (Huynh et al., 2021; Liang et al., 2020). As a result, before investing, a strategic assessment of market circumstances is essential, taking into account systemic risks such as Russia-Ukraine War induced economic sanctions sentiments.

## 4.3 | Additional robustness results

### 4.3.1 | Impact of war, sanction, and anxiety on G7 fixed income securities

Our above analysis uses composite indicator RUWESsent to examine the effect of war and sanction public sentiment on G7 fixed income securities. However, the effect of war, sanction and anxiety from those events might be different. The influence of sentiments caused by the conflict and sanctions may differ, since each situation brings new obstacles and opportunities for fixed income securities. While the outbreak of violence may initially create a negative reaction in markets, the introduction of economic sanctions may have a different effect. Indeed, sanctions could reassure investors and lessen the harmful impact of the conflict. Thus, we use Russia-Ukraine War sentiment index (RUWESsent\_War), Russia-Ukraine War Sanction sentiment index (RUWESsent\_Sanction), and Russia-Ukraine War anxiety index (RUWESsent\_Anxiety) and the result is presented in Table 3. The result indicates RUWESsent\_War, RUWESsent\_Sanction, and RUWESsent\_Anxiety has negative impact on G7 fixed income securities. The coefficient indicates the effect is stronger for RUWESsent\_War and highest negative impact is

TABLE 3 Impact of war, sanction, and anxiety.

<b>Panel A: War</b>							
Variables	(1) Canada	(2) France	(3) Germany	(4) Italy	(5) Japan	(6) UK	(7) US
RUWES sent_War	-0.013** (0.006)	-0.022*** (0.009)	-0.025*** (0.009)	-0.018** (0.008)	-0.018** (0.008)	-0.044*** (0.016)	-0.009** (0.004)
Constant	0.006 (0.005)	0.014* (0.008)	0.012 (0.007)	0.011 (0.007)	0.010 (0.007)	0.036*** (0.014)	0.003 (0.003)
Observations	337	337	337	337	337	337	337
R-squared	0.017	0.020	0.020	0.014	0.015	0.023	0.020
Adjusted R-squared	0.014	0.017	0.017	0.012	0.012	0.020	0.017
<b>Panel B: Sanction</b>							
Variables	(1) Canada	(2) France	(3) Germany	(4) Italy	(5) Japan	(6) UK	(7) US
RUWESsent_Sanction	-0.010*** (0.002)	-0.018*** (0.004)	-0.016*** (0.004)	-0.015*** (0.004)	-0.014*** (0.004)	-0.030*** (0.007)	-0.007*** (0.002)
Constant	0.009* (0.005)	0.019** (0.008)	0.016** (0.007)	0.015** (0.007)	0.014** (0.007)	0.043*** (0.014)	0.005 (0.003)
Observations	337	337	337	337	337	337	337
R-squared	0.054	0.054	0.054	0.046	0.047	0.053	0.054
Adjusted R-squared	0.051	0.052	0.051	0.044	0.045	0.050	0.052
<b>Panel C: Anxiety</b>							
Variables	(1) Canada	(2) France	(3) Germany	(4) Italy	(5) Japan	(6) UK	(7) US
RUWESsent_Anxiety	-0.007*** (0.002)	-0.012*** (0.004)	-0.011*** (0.004)	-0.009*** (0.004)	-0.009*** (0.004)	-0.021*** (0.007)	-0.004*** (0.002)
Constant	0.007 (0.005)	0.016* (0.008)	0.014* (0.008)	0.012* (0.007)	0.012* (0.007)	0.039*** (0.014)	0.004 (0.003)
Observations	337	337	337	337	337	337	337
R-squared	0.026	0.026	0.026	0.020	0.020	0.027	0.024
Adjusted R-squared	0.023	0.023	0.023	0.017	0.017	0.024	0.021

Note: This table presents the results of the impact of war, sanction, and anxiety on G7 fixed income securities. Panels A, B, and C shows the findings for war, sanction, and anxiety as independent variable using respective G7 fixed income securities return as dependent variable. All models are estimated using the Newey–West OLS estimator. Robust standard errors are in parentheses. \*\*\*, \*\*, \* denote significance at 1%, 5%, and 10% significance level.

witnessed on Germany. This result can be explained by the dependency of Germany on Russian energy, due to sanction there was interruption in the energy supply in Germany making energy costing higher and reflecting negative effect in the overall economy of Germany.<sup>7</sup> Our findings suggest economic sanctions and war sentiments signal to the markets that the international community is taking action to address the conflict, which uncertainty and affecting fixed income securities price, leading increased investors' loss aversion and herding behaviour. Previous studies (e.g., Bounou & Yatié, 2022; Yousaf et al., 2022) also discover the negative impact of Russia Ukraine war.

#### 4.3.2 | Robustness test using alternative proxy

The above analysis is based on the data of G7 treasury/sovereign/quasi-government bond indices. However, the effect of war and sanction public sentiment on G7 corporate bond might be different than the treasury/sovereign/quasi-government bond. This is because corporate bonds are issued by private companies, rather than governments, and are therefore subject to different economic and market forces. Furthermore, corporate bonds are more vulnerable to shifts in investor opinion than sovereign or quasi-government bonds. This is due to the fact

TABLE 4 Robustness test using corporate bond data.

<b>Panel A: RUWESsent</b>					
<b>Variables</b>	<b>(1) US_IGCB</b>	<b>(2) UK_IGCB</b>	<b>(3) Japan_IGCB</b>	<b>(4) EUROZONE_IGCB</b>	<b>(5) Canada_IGCB</b>
RUWESsent	-0.018*** (0.004)	-0.032*** (0.008)	-0.020*** (0.005)	-0.020*** (0.005)	-0.013*** (0.003)
Constant	0.007 (0.005)	0.020** (0.009)	0.009* (0.005)	0.010* (0.006)	0.005 (0.004)
Observations	337	337	337	337	337
R-squared	0.060	0.049	0.052	0.049	0.047
Adjusted R-squared	0.057	0.047	0.049	0.046	0.044
<b>Panel B: War</b>					
<b>Variables</b>	<b>(1) US_IGCB</b>	<b>(2) UK_IGCB</b>	<b>(3) Japan_IGCB</b>	<b>(4) EUROZONE_IGCB</b>	<b>(5) Canada_IGCB</b>
RUWESsent_War	-0.011** (0.005)	-0.019* (0.011)	-0.011* (0.006)	-0.011 (0.007)	-0.007 (0.005)
Constant	0.005 (0.005)	0.016* (0.009)	0.006 (0.006)	0.007 (0.006)	0.003 (0.004)
Observations	337	337	337	337	337
R-squared	0.012	0.010	0.009	0.008	0.007
Adjusted R-squared	0.009	0.007	0.006	0.005	0.004
<b>Panel C: Sanction</b>					
<b>Variables</b>	<b>(1) US_IGCB</b>	<b>(2) UK_IGCB</b>	<b>(3) Japan_IGCB</b>	<b>(4) EUROZONE_IGCB</b>	<b>(5) Canada_IGCB</b>
RUWESsent_Sanction	-0.009*** (0.002)	-0.016*** (0.005)	-0.010*** (0.003)	-0.010*** (0.003)	-0.007*** (0.002)
Constant	0.008 (0.005)	0.021** (0.009)	0.009* (0.006)	0.010* (0.006)	0.006 (0.004)
Observations	337	337	337	337	337
R-squared	0.044	0.036	0.040	0.036	0.035
Adjusted R-squared	0.041	0.033	0.037	0.033	0.032
<b>Panel D: Anxiety</b>					
<b>Variables</b>	<b>(1) US_IGCB</b>	<b>(2) UK_IGCB</b>	<b>(3) Japan_IGCB</b>	<b>(4) EUROZONE_IGCB</b>	<b>(5) Canada_IGCB</b>
RUWESsent_Anxiety	-0.006** (0.002)	-0.010** (0.005)	-0.006** (0.003)	-0.006** (0.003)	-0.004** (0.002)
Constant	0.006 (0.005)	0.018* (0.009)	0.007 (0.006)	0.008 (0.006)	0.004 (0.004)
Observations	337	337	337	337	337
R-squared	0.016	0.014	0.015	0.012	0.012
Adjusted R-squared	0.013	0.011	0.012	0.009	0.009

Note: This table presents the results of the impact of war, sanction, and anxiety on G7 corporate bonds. Panels A, B, C, and D shows the findings for composite indicator, war, sanction, and anxiety respectively as independent variable using respective G7 corporate bonds indices return as dependent variable. All models are estimated using the Newey–West OLS estimator. Robust standard errors are in parentheses. \*\*\*, \*\*, \* denote significance at 1%, 5%, and 10% significance level.

that corporate bonds are generally seen as riskier investments than government bonds, as they are dependent on the financial success of the underlying companies. Therefore, we use S&P 500 investment grade corporate bond index (US\_IGCB), S&P U.K. investment grade corporate bond index (UK\_IGCB), S&P Japan investment grade corporate bond index (Japan\_IGCB), S&P EUROZONE investment grade corporate bond index (EUROZONE\_IGCB), and S&P Canada investment grade corporate bond index

(Canada\_IGCB) to proxy the G7 corporate bonds. The result of robustness test using corporate bond data is presented in Table 4. The findings show there is a negative impact of RUWESsent, RUWESsent\_War, RUWESsent\_Sanction, and RUWESsent\_Anxiety on G7 corporate bond returns. This result further corroborates our baseline findings and indicates there is a negative effect of war and sanction public sentiment on G7 fixed income securities. To further examine the robustness of our findings, we use Corporate Bond

Market Distress Index (CBMDI) as the proxy of fixed income market turbulence and sentiment indicator. The CMDI is an indicator that measures the level of distress in the corporate bond market (Boyarchenko et al., 2022). The CMDI can impact bond market returns by influencing investor sentiment. When the CMDI rises, it can signal to investors that the corporate bond market is facing greater risks and uncertainties, which can lead to increased fear and caution, driving investors to shift their portfolios towards other assets. The result of the additional robustness test using CMDI is presented in Table A3. The findings also show there is a negative impact of CMDI on G7 corporate bonds.

Overall, our additional robustness test corroborates our baseline results and suggests there is a negative effect of war and sanction public sentiment on G7 fixed income securities.

#### 4.3.3 | Contagion effect of Russia-Ukraine war on G7 fixed income market

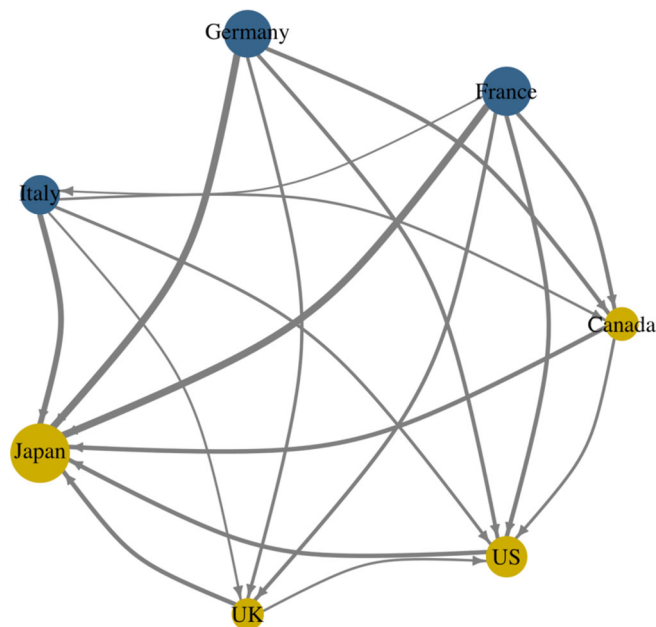
The above analysis clearly shows there is a negative effect of war and sanction public sentiment on G7 fixed income securities. However, the contagion effect of Russia-Ukraine war on G7 fixed income market and how war and sanction public sentiment drives or reduces the contagion is still unclear. Thus, we further examine the contagion effect of Russia-Ukraine war on G7 fixed income market using TVP-VAR based connectedness approach based on Antonakakis et al. (2020). For contagion analysis, we examine the volatility spillover across G7 fixed income market (Table 5).

First, we estimate the volatility of each indices using Univariate-Threshold GARCH model and estimate the static volatility connectedness using TVP-VAR model. The result of static volatility connectedness is presented in Table 4. The interaction between network variables is depicted in off-diagonal components, whereas own-variance shocks are presented in the on-diagonal. For example, the result of on-diagonal element shows United States has the highest own variance shock (37.67%), whereas France has the lowest own variance shock (21.02%). This finding suggests that 37.67% of the forecast error variance in the United States can be attributed to shocks within its own asset class. The row labelled “TO” indicates that France (94.50%) exhibits the highest spillover to other assets in the network, while the row labelled “FROM” shows that France (77.98%) experiences the highest spillover from other assets in the network. Based on the negative or positive values of NET, assets are classified as net risk receivers or transmitters. When considering typical net recipients and transmitters, France (16.52%) emerges as the primary net transmitter within the network, while Japan (−15.75%) appears to be the main net recipient. To gain a better understanding of the risk transmission mechanism, the results are visualised in a network connectedness plot depicted in Figure 5. Arrows in the figure indicate the direction of transmission, colour represents the nature of transmission (blue for net transmitters and yellow for net receivers), and the node size indicates the strength of risk transmission. The findings also show Japan is major receiver of risk and France is the major transmitter of risk. Overall, Japan, United Kingdom, United States and UK are in the net risk receiver, whereas Italy, Germany France are the major transmitter of shock. The total

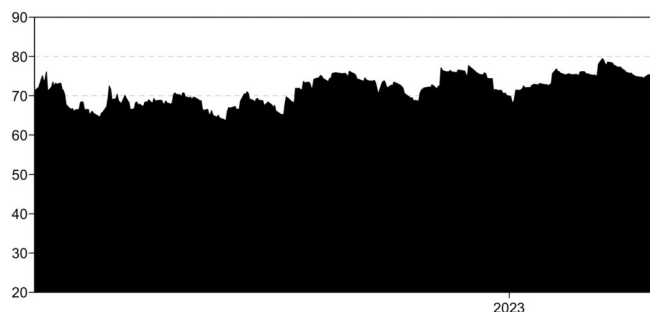
TABLE 5 Static connectedness results.

	Canada	France	Germany	Italy	Japan	UK	US	FROM
Canada	30.55	13.25	13.46	12.53	6.03	11.96	12.21	69.45
France	10.16	22.02	21.34	19.06	6.81	13.19	7.41	77.98
Germany	10.29	21.58	22.1	18.34	7.2	13.23	7.26	77.9
Italy	10.49	20.54	19.57	24.04	6.58	11.99	6.78	75.96
Japan	9.48	12.31	12.67	10.94	35.75	9.94	8.92	64.25
UK	11.56	16.09	15.75	13.82	6.55	26.09	10.14	73.91
US	14.53	10.72	10.24	9.22	5.58	12.05	37.67	62.33
TO	66.51	94.5	93.03	83.89	38.76	72.36	52.72	501.78
Inc. Own	97.06	116.52	115.13	107.93	74.51	98.46	90.39	TCI
NET	−2.94	16.52	15.13	7.93	−25.49	−1.54	−9.61	71.68

Note: Results are based on a generalised forecast error variance decomposition with 10 steps in advance and a 22-day rolling-window TVP-VAR model with a lag length of order 1 (BIC). TO = spillover to other asset; FROM = spillover from other asset; NET = net directional spillover, TCI = total connectedness index.



**FIGURE 5** Connectedness network plot. Results are based on a generalised forecast error variance decomposition with 10 steps in advance and a 22-day rolling-window TVP-VAR model with a lag length of order 1 (BIC). [Colour figure can be viewed at [wileyonlinelibrary.com](http://wileyonlinelibrary.com)]



**FIGURE 6** Dynamic total connectedness index.

connectedness index (TCI) shows 71.68% connectedness in the network, indicating considerable contagion across the G7 fixed income markets. Thus, result indicates there is contagion effect from Eurozone's to other G7 countries during Russia Ukraine war. However, the result is static and failed to uncover time-varying contagion effect, thus, we further examine dynamic connectedness.

The result of dynamic connectedness is presented in Figure 6. The result shows there is a time varying connectedness among G7 fixed income securities. The findings clearly show there is an upsurge in connectedness after the Russian invasion of Ukraine, indicating contagion effect of Russia Ukraine war on G7 fixed income securities. The result of net directional connectedness and pairwise connectedness is presented in Figures A2 and A3, and both figure shows there is contagion effect of Russia Ukraine war on G7 fixed income securities. The net-pairwise connectedness result shows contagion effect from Eurozone's to other G7 countries during Russia Ukraine war.

Finally, to examine the impact of war and sanction sentiment on G7 fixed income securities contagion, we regress TCI with the war and sanction sentiment indices and results are presented in Table 6. The findings indicate there is a positive impact of RUWESsent, RUWESsent\_War, RUWESsent\_Sanction, and RUWESsent\_Anxiety on G7 fixed income securities contagion. This result implies war and sanction sentiment drives contagion from European Markets to other G7 markets. Thus, during times of war and sanctions, negative sentiment in European markets can have a spillover effect on other G7 markets. This means that investor sentiment towards the conflict and sanctions can drive changes in market performance and affect other markets beyond just Europe. Our finding is consistent with earlier studies who documented

**TABLE 6** Impact of war sentiment, sanction sentiment, and anxiety on contagion across G7 debt market.

Variables	(1) TCI	(2) TCI	(3) TCI	(4) TCI
RUWESsent	0.246*** (0.051)			
RUWESsent_War		0.246*** (0.050)		
RUWESsent_Sanction			0.212*** (0.038)	
RUWESsent_Anxiety				0.215*** (0.044)
Constant	68.876*** (0.564)	68.264*** (0.671)	69.598*** (0.387)	69.859*** (0.379)
Observations	338	338	338	338
R-squared	0.372	0.367	0.293	0.287
Adjusted R-squared	0.370	0.365	0.291	0.284

Note: This table presents the results of impact of war sentiment, sanction sentiment, and anxiety on contagion across G7 debt market. Model 1, 2, 3, and 4 is developed using composite indicator, war, sanction, and anxiety respectively as independent variable. All models are estimated using the Newey–West OLS estimator. Robust standard errors are in parentheses. \*\*\*, \*\*, \* denote significance at 1%, 5%, and 10% significance level.

negative influence of Russia-Ukraine war on other financial markets (Ahmed et al., 2023; Boubaker et al., 2022; Bounougou & Yatié, 2022; Costola & Lorusso, 2022; Yousaf et al., 2022). Overall, findings suggest investors must carefully consider the potential contagion effects of war and sanctions when making investment decisions.

## 5 | CONCLUSION

Given the increasing role of social media platforms in exchanging information and ideas among investors, academicians and researchers in recent years have implemented many investor sentiment proxies based on news shared on social media outlets. We use behavioural indicators across social media, news media, and internet attention to reflect the public sentiment and examine the implications of war and sanctions sentiment for the G7 debt market during the Russia-Ukraine war. We apply the QQR and RWWC methods. This allows differentiating between correlation patterns in the upper, median, and lower quantiles. The RWWC is used to provide evidence on the dynamics of the relationship between war- and sanctions-related sentiments and the G7 debt market. We find a strong negative correlation between public sentiments and fixed income securities returns at the start of the invasion and during major invasion events, implying that war- and sanctions-induced public sentiments had a time-varying influence. Furthermore, we find evidence of negative correlation in the higher decomposition level during the financial turbulence period. Our additional analysis uncovers the robustness of our findings using alternatives proxies and our contagion analysis shows significant influences of war and sanction sentiment on spillover.

The findings of this study have significant implications. For portfolio investors, the time-varying relationship observed between our sentiment index and the G7 debt market can assist market participants in adopting more effective hedging strategies and portfolio diversification techniques, ultimately leading to improved returns during periods of market volatility. In summary, our research holds substantial importance in the formulation of asset allocation strategies, portfolio performance enhancement, and risk hedging. It should be noted that the assumption of homogeneity among market participants and economic agents lacks empirical evidence. Therefore, any analysis exploring the relationship between sentiments and financial assets should consider the possibility that economic agents are not homogeneous. From a policy perspective, policymakers can benefit from understanding whether a

strong dependency exists between investor sentiments and G7 debt, particularly under extreme negative shocks. This understanding can guide decision-making regarding the implementation of specific policies aimed at protecting investors from severe fluctuations in the financial market, especially during times of wars and war-related events. Lastly, this research can contribute to the development of policies aimed at mitigating the financial impact of the Russia-Ukraine war on the transmission of shocks between financial markets.

Our study is not without limitations. While our focus is on the G7 debt market, it would be valuable to expand the analysis to encompass other financial markets and asset classes. Doing so would provide a more comprehensive understanding of the impact of sentiments across various domains.

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## DATA AVAILABILITY STATEMENT

The data that support the findings of this study are available from the corresponding author upon reasonable request.

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## ENDNOTES

- <sup>1</sup> For example, Bounougou and Yatié (2022) and Khalfaoui et al. (2023) use Wikipedia trends and Google trends, respectively, to measure public attention towards the Russia-Ukraine war.
- <sup>2</sup> <https://www.cfainstitute.org/en/membership/professional-development/refresher-readings/fixed-income-markets-issuance-trading-funding>
- <sup>3</sup> <https://www.sipri.org/news/2023/world-military-expenditure-reaches-new-record-high-european-spending-surges-0>
- <sup>4</sup> <https://www.cgdev.org/publication/ukrainian-war-and-feds-interest-rate-hikes-double-whammy-emerging-markets-and>
- <sup>5</sup> We use S&P Dow Jones indices to proxy the G7 fixed income securities. See <https://www.spglobal.com/spdji/en/index-family/fixed-income/treasury-sovereign-quasi-government/#overview>.
- <sup>6</sup> See Troster (2018) for detailed methodology of quantile granger causality test and Antonakakis et al. (2020) for TVP-VAR based connectedness approach.
- <sup>7</sup> Source: <https://www.dw.com/en/ukraine-war-costs-germanys-economy-100-billion/a-64768176>

## REFERENCES

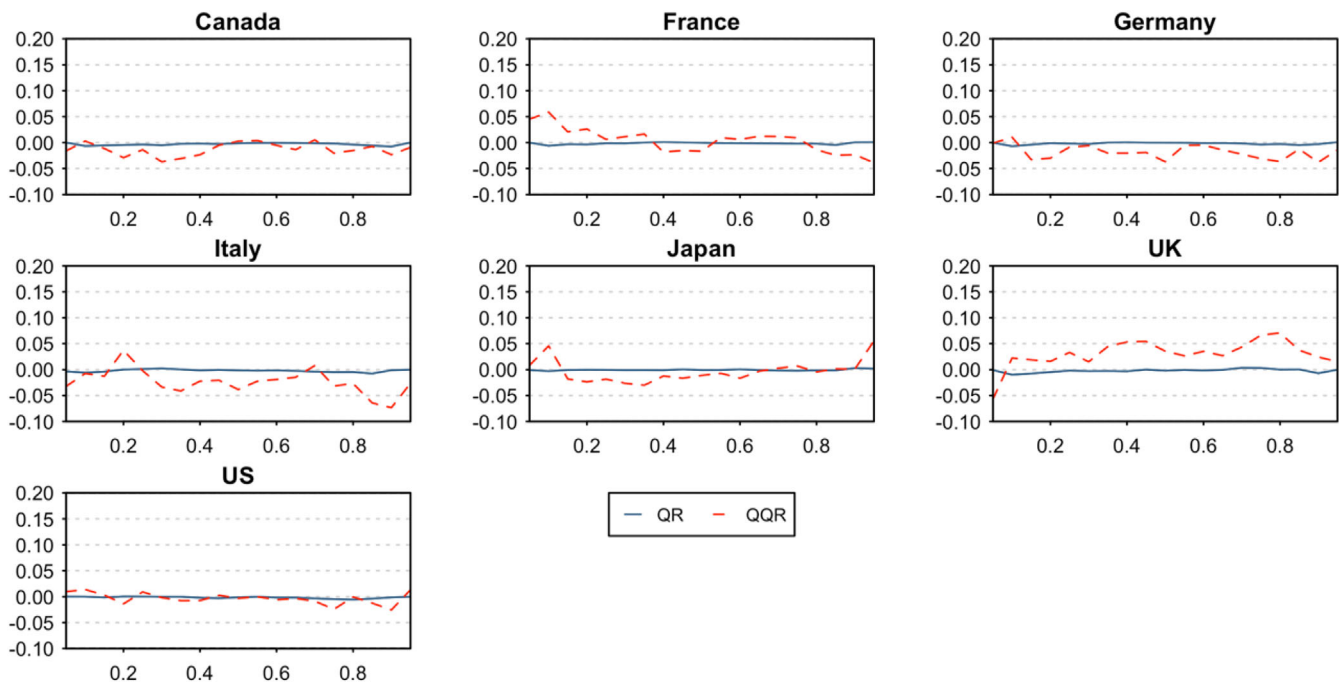
- Abakah, E. J. A., Adeabah, D., Tiwari, A. K., & Abdullah, M. (2022). Analyzing the effect of public sentiment towards economic sanctions news during Russia-Ukraine conflict on blockchain market and fintech industry. Available at SSRN 4359071.
- Adekoya, O. B., Oliyide, J. A., Yaya, O. S., & Al-Faryan, M. A. S. (2022). Does oil connect differently with prominent assets during war? Analysis of intra-day data during the Russia-Ukraine saga. *Resources Policy*, 77, 102728. <https://doi.org/10.1016/j.resourpol.2022.102728>
- Ahmed, S., Hasan, M. M., & Kamal, M. R. (2023). Russia-Ukraine crisis: The effects on the European stock market. *European Financial Management*, 29(4), 1078–1118. <https://doi.org/10.1111/eufm.12386>
- Antonakakis, N., Chatziantoniou, I., & Gabauer, D. (2020). Refined measures of dynamic connectedness based on time-varying parameter vector autoregressions. *Journal of Risk and Financial Management*, 13(4), 84. <https://www.mdpi.com/1911-8074/13/4/84>
- Aslanidis, N., Bariviera, A. F., & López, Ó. G. (2022). The link between cryptocurrencies and Google Trends attention. *Finance Research Letters*, 47, 102654. <https://doi.org/10.1016/j.frl.2021.102654>
- Baur, D. G., & Lucey, B. M. (2010). Is gold a hedge or a safe haven? An analysis of stocks, bonds and gold. *Financial Review*, 45(2), 217–229. <https://doi.org/10.1111/j.1540-6288.2010.00244.x>
- Będowska-Sójka, B., Demir, E., & Zaremba, A. (2022). Hedging geopolitical risks with different asset classes: A focus on the Russian invasion of Ukraine. *Finance Research Letters*, 50, 103192. <https://doi.org/10.1016/j.frl.2022.103192>
- Bosman, A., Gubareva, M., & Teplova, T. (2023). Asymmetric effects of geopolitical risk on major currencies: Russia-Ukraine tensions. *Finance Research Letters*, 51, 103440. <https://doi.org/10.1016/j.frl.2022.103440>
- Boubaker, S., Goodell, J. W., Pandey, D. K., & Kumari, V. (2022). Heterogeneous impacts of wars on global equity markets: Evidence from the invasion of Ukraine. *Finance Research Letters*, 48, 102934. <https://doi.org/10.1016/j.frl.2022.102934>
- Boungou, W., & Yatié, A. (2022). The impact of the Ukraine-Russia war on world stock market returns. *Economics Letters*, 215, 110516. <https://doi.org/10.1016/j.econlet.2022.110516>
- Boyarchenko, N., Crump, R., Kovner, A., & Shachar, O. (2022). *What is corporate bond market distress?* Federal Reserve Bank of New York <https://libertystreeteconomics.newyorkfed.org/2022/06/what-is-corporate-bond-market-distress/>
- Caldara, D., & Iacoviello, M. (2022). Measuring Geopolitical Risk. *American Economic Review*, 112(4), 1194–1225. <https://doi.org/10.1257/aer.20191823>
- Chortane, S. G., & Pandey, D. K. (2022). Does the Russia-Ukraine war lead to currency asymmetries? A US dollar tale. *The Journal of Economic Asymmetries*, 26, e00265. <https://doi.org/10.1016/j.jeca.2022.e00265>
- Choudhry, T. (2010). World War II events and the Dow Jones industrial index. *Journal of Banking & Finance*, 34(5), 1022–1031. <https://doi.org/10.1016/j.jbankfin.2009.11.004>
- Costantini, M., & Sousa, R. M. (2022). What uncertainty does to euro area sovereign bond markets: Flight to safety and flight to quality. *Journal of International Money and Finance*, 122, 102574. <https://doi.org/10.1016/j.jimonfin.2021.102574>
- Costola, M., & Lorusso, M. (2022). Spillovers among energy commodities and the Russian stock market. *Journal of Commodity Markets*, 22, 100249. <https://doi.org/10.1016/j.jcomm.2022.100249>
- Daubechies, I. (1992). *Ten lectures on wavelets*. SIAM.
- Diaconășu, D. E., Mehdian, S. M., & Stoica, O. (2022). The reaction of financial markets to Russia's invasion of Ukraine: Evidence from gold, oil, bitcoin, and major stock markets. *Applied Economics Letters*, 1–5. <https://doi.org/10.1080/13504851.2022.2107608>
- Feng, C., Han, L., Vigne, S., & Xu, Y. (2023). Geopolitical risk and the dynamics of international capital flows. *Journal of International Financial Markets, Institutions and Money*, 82, 101693. <https://doi.org/10.1016/j.intfin.2022.101693>
- Frey, B. S., & Kucher, M. (2000). World War II as reflected on capital markets. *Economics Letters*, 69(2), 187–191. [https://doi.org/10.1016/S0165-1765\(00\)00269-X](https://doi.org/10.1016/S0165-1765(00)00269-X)
- Frey, B. S., & Kucher, M. (2001). Wars and markets: How bond values reflect the Second World War. *Economica*, 68(271), 317–333. <https://doi.org/10.1111/1468-0335.00249>
- Gençay, R., Selçuk, F., & Whitcher, B. J. (2001). *An introduction to wavelets and other filtering methods in finance and economics*. Elsevier.
- Halousková, M., Stašek, D., & Horváth, M. (2022). The role of investor attention in global asset price variation during the invasion of Ukraine. *Finance Research Letters*, 50, 103292.
- Hudson, R., & Urquhart, A. (2015). War and stock markets: The effect of World War Two on the British stock market. *International Review of Financial Analysis*, 40, 166–177. <https://doi.org/10.1016/j.irfa.2015.05.015>
- Huynh, T. L. D., Foglia, M., Nasir, M. A., & Angelini, E. (2021). Feverish sentiment and global equity markets during the COVID-19 pandemic. *Journal of Economic Behavior & Organization*, 188, 1088–1108. <https://doi.org/10.1016/j.jebo.2021.06.016>
- Izzeldin, M., Muradoğlu, Y. G., Pappas, V., Petropoulou, A., & Sivaprasad, S. (2023). The impact of the Russian-Ukrainian war on global financial markets. *International Review of Financial Analysis*, 87, 102598. <https://doi.org/10.1016/j.irfa.2023.102598>
- Karkowska, R., & Urjasz, S. (2023). How does the Russian-Ukrainian war change connectedness and hedging opportunities? Comparison between dirty and clean energy markets versus global stock indices. *Journal of International Financial Markets, Institutions and Money*, 85, 101768. <https://doi.org/10.1016/j.intfin.2023.101768>
- Khalfaoui, R., Gozgor, G., & Goodell, J. W. (2023). Impact of Russia-Ukraine war attention on cryptocurrency: Evidence from quantile dependence analysis. *Finance Research Letters*, 52, 103365. <https://doi.org/10.1016/j.frl.2022.103365>
- Kumari, V., Kumar, G., & Pandey, D. K. (2022). Are the European Union stock markets vulnerable to the Russia-Ukraine War? Available at SSRN 100793.
- Kumari, V., Kumar, G., & Pandey, D. K. (2023). Are the European Union stock markets vulnerable to the Russia-Ukraine war? *Journal of Behavioral and Experimental Finance*, 37, 100793. <https://doi.org/10.1016/j.jbef.2023.100793>
- Leippold, M., & Matthys, F. (2022). Economic policy uncertainty and the yield curve. *Review of Finance*, 26(4), 751–797. <https://doi.org/10.1093/rof/rfac031>
- Liang, C., Tang, L., Li, Y., & Wei, Y. (2020). Which sentiment index is more informative to forecast stock market volatility?

- Evidence from China. *International Review of Financial Analysis*, 71, 101552. <https://doi.org/10.1016/j.irfa.2020.101552>
- Lo, G.-D., Marcellin, I., Bassène, T., & Sène, B. (2022). The Russo-Ukrainian war and financial markets: the role of dependence on Russian commodities. *Finance Research Letters*, 50, 103194. <https://doi.org/10.1016/j.frl.2022.103194>
- Long, H., Demir, E., Będowska-Sójka, B., Zaremba, A., & Shahzad, S. J. H. (2022). Is geopolitical risk priced in the cross-section of cryptocurrency returns? *Finance Research Letters*, 49, 103131. <https://doi.org/10.1016/j.frl.2022.103131>
- Long, J. B. D., Shleifer, A., Summers, L. H., & Waldmann, R. J. (1990). Noise trader risk in financial markets. *Journal of Political Economy*, 98(4), 703–738. <https://doi.org/10.1086/261703>
- Mohamad, A. (2022). Safe flight to which haven when Russia invades Ukraine? A 48-hour story. *Economics Letters*, 216, 110558. <https://doi.org/10.1016/j.econlet.2022.110558>
- Müller, U. A., Dacorogna, M. M., Davé, R. D., Pictet, O. V., Olsen, R. B., & Ward, J. R. (1993). *Fractals and intrinsic time: A challenge to econometricians*. Olsen & Associates.
- Polanco-Martínez, J. M., Fernández-Macho, J., Neumann, M. B., & Faria, S. H. (2018). A pre-crisis vs. crisis analysis of peripheral EU stock markets by means of wavelet transform and a nonlinear causality test. *Physica A: Statistical Mechanics and its Applications*, 490, 1211–1227. <https://doi.org/10.1016/j.physa.2017.08.065>
- Qureshi, A., Rizwan, M. S., Ahmad, G., & Ashraf, D. (2022). Russia-Ukraine war and systemic risk: Who is taking the heat? *Finance Research Letters*, 48, 103036. <https://doi.org/10.1016/j.frl.2022.103036>
- Rehman, M. U. (2020). Dynamic correlation pattern amongst alternative energy market for diversification opportunities. *Journal of Economic Structures*, 9(1), 16. <https://doi.org/10.1186/s40008-020-00197-2>
- Schneider, G., & Troeger, V. E. (2006). War and the world economy: stock market reactions to international conflicts. *Journal of Conflict Resolution*, 50(5), 623–645. <https://doi.org/10.1177/0022002706290430>
- Sim, N., & Zhou, H. (2015). Oil prices, US stock return, and the dependence between their quantiles. *Journal of Banking & Finance*, 55, 1–8. <https://doi.org/10.1016/j.jbankfin.2015.01.013>
- Tosun, O. K., & Eshraghi, A. (2022). Corporate decisions in times of war: Evidence from the Russia-Ukraine conflict. *Finance Research Letters*, 48, 102920. <https://doi.org/10.1016/j.frl.2022.102920>
- Troster, V. (2018). Testing for Granger-causality in quantiles. *Econometric Reviews*, 37(8), 850–866. <https://doi.org/10.1080/07474938.2016.1172400>
- Umar, M., Riaz, Y., & Yousaf, I. (2022). Impact of Russian-Ukraine war on clean energy, conventional energy, and metal markets: Evidence from event study approach. *Resources Policy*, 79, 102966. <https://doi.org/10.1016/j.resourpol.2022.102966>
- Umar, Z., Polat, O., Choi, S.-Y., & Teplova, T. (2022). The impact of the Russia-Ukraine conflict on the connectedness of financial markets. *Finance Research Letters*, 48, 102976. <https://doi.org/10.1016/j.frl.2022.102976>
- Verdickt, G. (2020). The effect of war risk on managerial and investor behavior: Evidence from the brussels stock exchange in the Pre-1914 Era. *The Journal of Economic History*, 80(3), 629–669. <https://doi.org/10.1017/S0022050720000303>
- Wang, A. Y., & Young, M. (2020). Terrorist attacks and investor risk preference: Evidence from mutual fund flows. *Journal of Financial Economics*, 137(2), 491–514. <https://doi.org/10.1016/j.jfineco.2020.02.008>
- Yousaf, I., Patel, R., & Yarovaya, L. (2022). The reaction of G20+ stock markets to the Russia-Ukraine conflict.
- Zaremba, A., Kizys, R., Aharon, D. Y., & Umar, Z. (2022). Term spreads and the COVID-19 pandemic: Evidence from international sovereign bond markets. *Finance Research Letters*, 44, 102042. <https://doi.org/10.1016/j.frl.2021.102042>
- Zhang, W., Wang, P., Li, X., & Shen, D. (2018). Twitter's daily happiness sentiment and international stock returns: Evidence from linear and nonlinear causality tests. *Journal of Behavioral and Experimental Finance*, 18, 50–53. <https://doi.org/10.1016/j.jbef.2018.01.005>

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## APPENDIX

## QUANTILE ON QUANTILE REGRESSION ROBUSTNESS TEST

FIGURE A1 QQR results robustness test. [Colour figure can be viewed at [wileyonlinelibrary.com](https://onlinelibrary.wiley.com/doi/10.1002/jfe.2387)]

## QUANTILE GRANGER CAUSALITY TEST RESULTS

TABLE A1 Quantile granger test results of RUWESsent to fixed income assets.

$\tau$	Canada	France	Germany	Italy	Japan	UK	US
0.05	1.21	1.18	1.15	1.27	1.57*	1.23	1.43*
0.10	1.74**	1.72**	1.79**	1.74**	2.17**	2.26**	2.02**
0.15	1.49*	1.37*	1.28	1.97**	1.48*	2.27**	1.65**
0.20	1.51*	1.43*	1.44*	1.64*	1.83**	2.12**	1.52*
0.25	1.8**	1.42*	1.43*	1.92**	1.61*	2.48***	1.90**
0.30	1.78**	1.83**	1.90**	2.18**	2.15**	2.83***	1.72**
0.35	1.87**	2.26**	2.43***	2.53***	2.35***	3.37***	1.93**
0.40	1.87**	2.43***	2.32**	2.43***	2.09**	3.31***	1.79**
0.45	2.08**	2.10**	1.98**	2.05**	1.95**	2.79***	1.80**
0.50	1.82**	2.18**	2.11**	2.21**	2.11**	2.91***	1.93**
0.55	2.00**	2.52***	2.66***	2.60***	2.18**	3.05***	2.15**
0.60	1.97**	2.2**	2.39***	2.09**	1.99**	2.66***	2.09**
0.65	2.25**	2.44***	2.64***	1.95**	2.23**	2.75***	2.35***
0.70	1.95**	2.13**	2.30**	1.78**	1.81**	2.71***	2.26**
0.75	2.04**	2.19**	2.30**	1.58*	1.84**	2.39***	2.04**
0.80	1.79**	1.84**	1.92**	1.31*	1.08	1.92**	2.02**
0.85	1.50*	1.69**	1.67**	1.47*	1.16	1.50*	1.56*
0.90	1.23	1.19	1.30*	0.93	0.94	1.26	1.28
0.95	0.71	0.56	0.56	0.57	0.87	0.58	0.68

Note: \*\*\*, \*\*, and \* indicate significance at the 1%, 5% and 10% levels, respectively. Results are based on the estimation following quantile Granger casualty test (Troster, 2018).

$\tau$	Canada	France	Germany	Italy	Japan	UK	US
0.05	0.41	0.51	0.25	0.53	0.55	0.32	0.34
0.10	0.51	0.28	0.23	0.44	0.90	0.44	0.40
0.15	0.44	0.46	0.15	0.63	0.68	0.43	0.98
0.20	0.75	0.52	0.23	0.76	0.83	0.31	1.08
0.25	0.80	0.53	0.25	0.81	1.15	0.42	1.15
0.30	0.89	0.57	0.19	0.87	1.62*	0.35	0.94
0.35	0.89	0.53	0.12	0.74	1.5*	0.43	0.76
0.40	0.99	0.62	0.10	0.91	1.43*	0.52	0.74
0.45	1.11	0.44	0.13	0.87	1.46*	0.63	0.65
0.50	0.94	0.42	0.14	0.85	1.31*	0.78	0.88
0.55	0.99	0.59	0.21	0.76	1.4*	0.84	0.95
0.60	0.66	0.50	0.13	0.93	1.28	0.60	1.05
0.65	0.56	0.33	0.33	1.01	0.86	0.78	1.04
0.70	0.50	0.76	0.30	0.85	0.93	0.90	1.24
0.75	0.48	0.90	0.46	0.76	0.67	0.71	1.29*
0.80	0.41	0.80	0.45	0.51	0.58	0.77	1.15
0.85	0.71	0.43	0.26	0.53	0.44	0.52	0.99
0.90	0.63	0.30	0.13	0.39	0.38	0.86	0.59
0.95	0.41	0.30	0.02	0.17	0.30	0.30	0.72

TABLE A2 Quantile granger test results of fixed income assets to RUWESent.

Note: \*\*\*, \*\*, and \* indicate significance at the 1%, 5%, and 10% levels, respectively. Results are based on the estimation following quantile Granger casualty test (Troster, 2018).

CONTAGION ACROSS G7 DEBT MARKET RESULTS

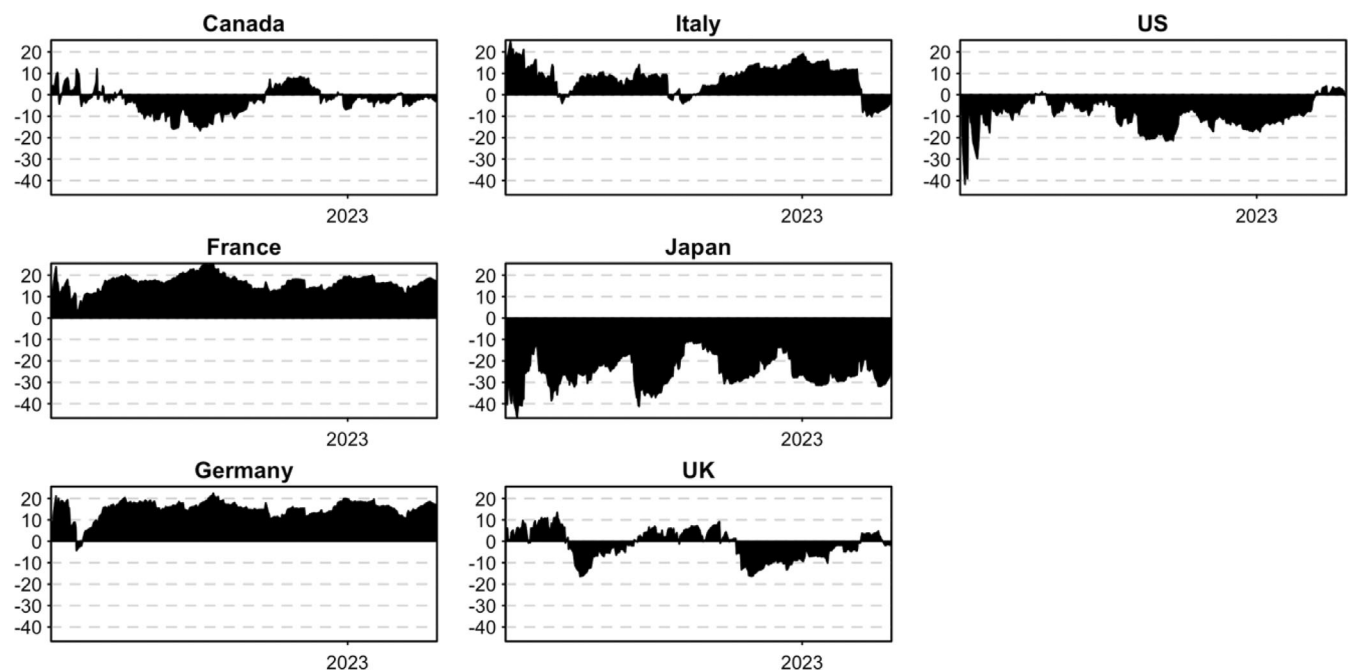


FIGURE A2 Dynamic net directional connectedness. Results are based on a generalised forecast error variance decomposition with 10 steps in advance and a 22-day rolling-window TVP-VAR model with a lag length of order 1 (BIC).

TABLE A3 Additional robustness test: Impact of CMDI on corporate bonds.

Variables	(1) US_IGCB	(2) UK_IGCB	(3) Japan_IGCB	(4) EUROZONE_IGCB	(5) Canada_IGCB
CMDI	-0.143*** (0.019)	-0.299*** (0.040)	-0.213*** (0.023)	-0.216*** (0.027)	-0.124*** (0.017)
Constant	0.010 (0.007)	0.028* (0.016)	0.016* (0.008)	0.017* (0.010)	0.008 (0.006)
Observations	66	66	66	66	66
R-squared	0.466	0.466	0.609	0.524	0.461
Adjusted R-squared	0.457	0.457	0.602	0.516	0.452

Note: This table presents the results of the impact of CMDI on G7 corporate bonds. All models are estimated using the Newey–West OLS estimator. Robust standard errors are in parentheses. \*\*\*, \*\*, \* denote significance at 1%, 5%, and 10% significance level.

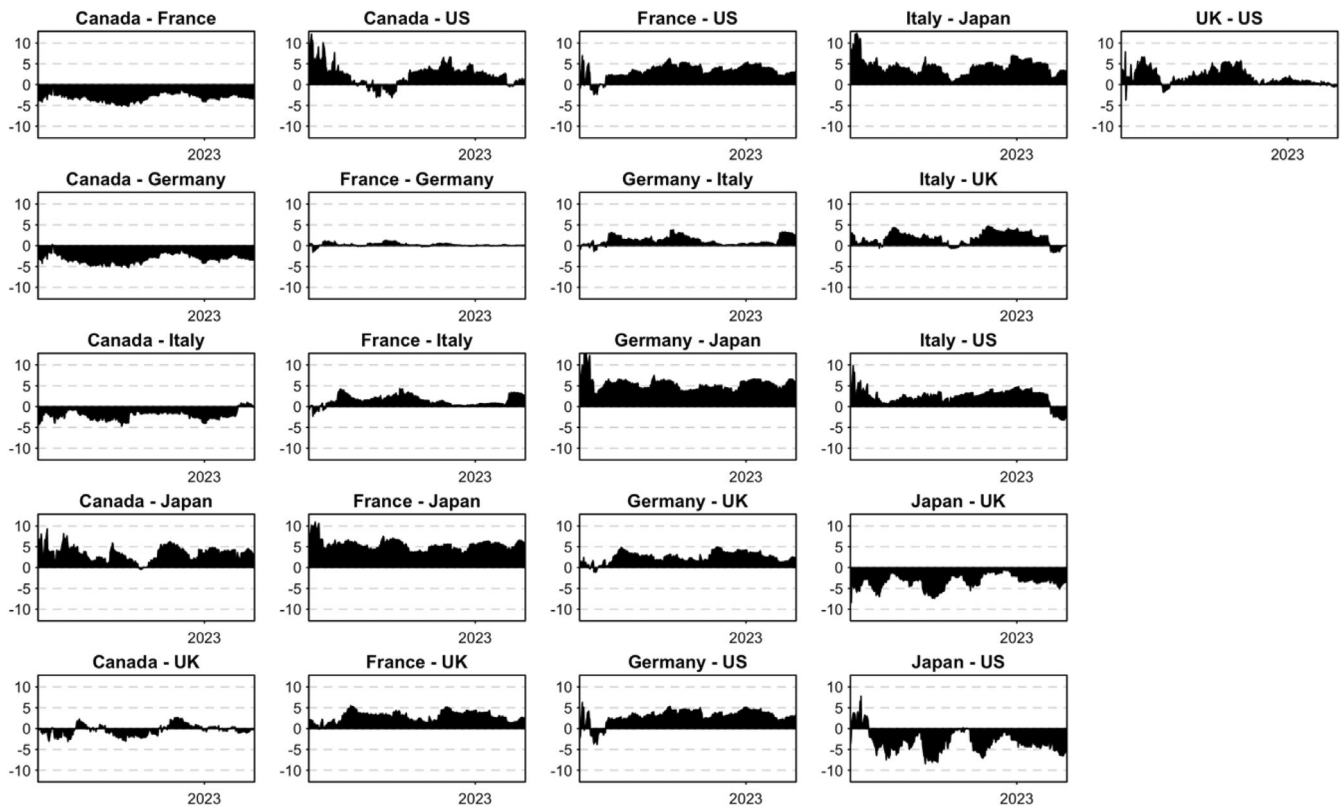


FIGURE A3 Dynamic net pairwise directional connectedness. Results are based on a generalised forecast error variance decomposition with 10 steps in advance and a 22-day rolling-window TVP-VAR model with a lag length of order 1 (BIC).